

Exam
Introduction Mathematical Statistics
Semster I 2024-2025

Family name:

First name:

Student number:

Remarks:

- The exam consists of **6** tasks.
- You have **180** minutes to complete the exam.
- You can only use the provided short version of our lecture notes, the handbook of distributions and the R introduction.
- The solution must be documented well with calculations and if necessary references to theorems of our lecture. Giving just the solution is not sufficient.
- Whenever you need a certain quantile or probability try to express it using quantiles respectively the distribution function of standard distributions, like standard normal, t or F distribution.

Task	1	2	3	4	5	6	Σ
Points possible	13	11	9	8	5	6	52
Points achieved							

Grade exam:

Grade homework:

Final grade:

Task 1 (3 + 3 + 1 + 1 + 2 + 3)

In this task (and the whole exam) you can use that for independent and identically distributed random variables $X_1, \dots, X_n$ with distribution function F :

$$F_{\min(X_1, \dots, X_n)}(x) = 1 - [1 - F(x)]^n.$$

Let $X_1, \dots, X_n$ be independent and identically distributed with density $f_a, a \in \mathbb{R}$ given by

$$f_a(x) = \begin{cases} e^{-(x-a)} & \text{if } x \geq a \\ 0 & \text{if } x < a. \end{cases}$$

a) Show that

$$\hat{a} := \min(X_1, \dots, X_n)$$

is the maximum likelihood estimator of a .

b) Show that the distribution function of $\hat{a}$ equals

$$F_{\hat{a}}(x) = \begin{cases} 1 - e^{n(a-x)} & x \geq a \\ 0 & x < a. \end{cases}$$

We define $T := \hat{a} - a$.

c) Show that

$$F_T(x) = \begin{cases} 1 - e^{-nx} & x \geq 0 \\ 0 & x < 0 \end{cases}.$$

d) Show that

$$F_T^{-1}(p) = -\frac{\log(1-p)}{n} \quad \text{for } p \in (0, 1).$$

e) Construct a two-sided $1 - \alpha$ confidence interval for a .

[You will not get points for the trivial confidence interval $(-\infty, \infty)$.]

f) Is $\hat{a}$ a (weakly) consistent estimator (for a)?

Task 2 (2+4+2+3)

Let $X_1, \dots, X_n$ be independent and identically distributed with a discrete density given by

$$f_a(x) = \begin{cases} \frac{1}{3} - a & x = -1 \\ \frac{1}{3} & x = 0 \\ \frac{1}{3} + a & x = 1 \end{cases}$$

for $-\frac{1}{3} \leq a \leq \frac{1}{3}$.

- a) Show that there exists a moment estimator defined by

$$\hat{a} = \frac{\overline{X}}{2}.$$

Note that this estimator can produce values smaller than $-\frac{1}{3}$ or larger than $\frac{1}{3}$. We will not worry about this unfavourable behaviour during the exam.

- b) Compute the MSE of $\hat{a}$.

- c) Show that

$$\sqrt{n}\sqrt{12} \left(\frac{\hat{a} - a}{\sqrt{2 - 12a^2}} \right) \xrightarrow{\mathcal{D}} N(0, 1).$$

- d) Construct an two-sided asymptotic $1 - \alpha$ confidence interval for a .

[You will not get points for the trivial confidence interval $(-1/3, 1/3)$.]

Task 3 (4+3+2)

Let $\begin{pmatrix} X_1 \\ Y_1 \end{pmatrix}, \dots, \begin{pmatrix} X_n \\ Y_n \end{pmatrix}$ be independent and identically distributed random vectors in $\mathbb{R}^2$ following a multivariate normal distribution $N(\mu, \Sigma)$.

- a) Show that $\begin{pmatrix} \overline{X} \\ \overline{Y} \end{pmatrix} \sim N(\mu, \frac{1}{n}\Sigma)$.

- b) Show that

$$n \left[\begin{pmatrix} \overline{X} \\ \overline{Y} \end{pmatrix} - \mu \right]' \Sigma^{-1} \left[\begin{pmatrix} \overline{X} \\ \overline{Y} \end{pmatrix} - \mu \right] \sim \chi_2^2.$$

Now we want to test

$$H_0 : \mu = \mu_0 \quad \text{vs.} \quad H_1 : \mu \neq \mu_0$$

if Σ is known.

- c) Construct a test of level α for this hypothesis.

Task 4 (3+3+2)

Let $Y = X\beta + \epsilon$ where $X = \begin{pmatrix} x_{1,1} & x_{1,2} \\ \vdots & \vdots \\ x_{n,1} & x_{n,2} \end{pmatrix}$ fulfils $\det(X'X) \neq 0$ and $\epsilon = \begin{pmatrix} \epsilon_1 \\ \vdots \\ \epsilon_n \end{pmatrix}$

where $\epsilon_1, \dots, \epsilon_n$ are independent and identically distributed with $\epsilon_1 \sim N(0, \sigma^2)$.

Further we denote by $\hat{\beta} = \begin{pmatrix} \hat{\beta}_1 \\ \hat{\beta}_2 \end{pmatrix}$ and $\hat{\sigma}^2$ the ML estimators of β and σ^2 .

- a) Show that $\hat{\beta}_1 - \hat{\beta}_2 \sim N(\beta_1 - \beta_2, Z)$ where $Z = \sigma^2(a - 2b + c)$ and a, b, c are defined by

$$\begin{pmatrix} a & b \\ b & c \end{pmatrix} = (X'X)^{-1}.$$

- b) Show that

$$\frac{\hat{\beta}_1 - \hat{\beta}_2 - (\beta_1 - \beta_2)}{\sqrt{\frac{n}{n-2}\hat{\sigma}^2(a - 2b + c)}} \sim t_{n-2}.$$

- c) We define a test by

$$\Psi(X, Y) = \begin{cases} 1 & \frac{\hat{\beta}_1 - \hat{\beta}_2}{\sqrt{\frac{n}{n-2}\hat{\sigma}^2(a - 2b + c)}} > F_{t_{n-2}}^{-1}(1 - \alpha) \\ 0 & \leq \end{cases}$$

where $F_{t_{n-2}}^{-1}(1 - \alpha)$ denotes the $1 - \alpha$ quantile of a t-distribution with $n - 2$ degrees of freedom. Show that this is a level α test for

$$H_0: \beta_1 \leq \beta_2 \text{ vs. } H_1: \beta_1 > \beta_2.$$

Task 5 (2 + 1 + 1 + 1)

The following R-function estimates a statistical model based on a data-vector X representing a sample $X_1, \dots, X_n$.

```
fu <- function(X) {
  A <- function(Y) {
    return(sum(log(dweibull(X,Y[1],Y[2]))))
  }
  B <- function(Y) return(-A(Y))
  C <- optim(c(1,1),B)$par
  D <- matrix(ncol=2,nrow=1000)
  for(i in 1:1000) {
    F <- sample(X,replace=TRUE)
    G <- function(Y) {
      return(-sum(log(dweibull(F,Y[1],Y[2]))))
    }
    D[i,] <- optim(c(1,1),G)$par
  }
  H <- sort(D[,1])[c(1+floor(1000*0.025),1000-floor(1000*0.025))]
  G <- sort(D[,2])[c(1+floor(1000*0.025),1000-floor(1000*0.025))]
  return(list(C,H,G))
}
```

- a) Which statistical model is assumed?
- b) What is represented by A?
- c) What values are represented by C?
- d) What values are represented by H?

Task 6 (3+3)

We want to check whether there is dependence between the average income of an inhabitant of a municipality and the voting result of the PVV at the last election. We have data from 207 municipalities.

Municipality	Average income in 1000 €	Voting share of PVV
Aalsmeer	31.3	25.1
Aalten	24.4	19.8
Achtkarspelen	22.0	34.5
$\vdots$	$\vdots$	$\vdots$
Zwijndrecht	26.5	29.6
Zwolle	27.3	16.1

The following pages contain some analysis in R. Use the information for the following tasks.

- a) Choose a suitable statistical model and justify your choice. Formulate a reasonable null- and alternative hypothesis.
- b) Choose a suitable test for your statistical model using a significance level of $\alpha = 0.05$. Can you reject the null-hypothesis? What can you conclude?

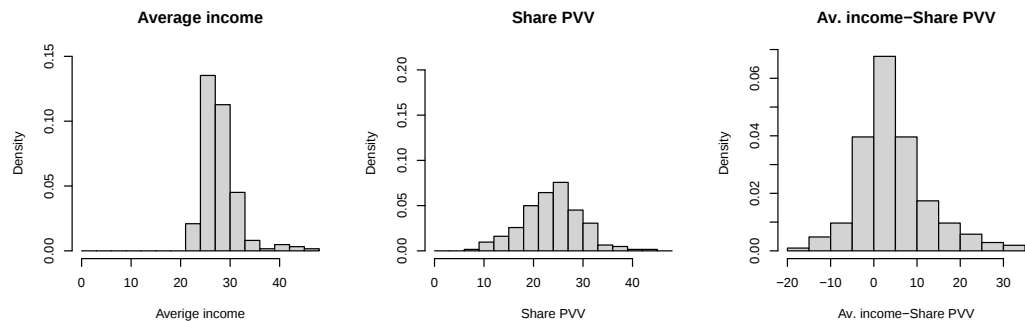


Figure 1: Histograms of the variables: average income (left), voting share PVV (middle), average income - voting share PVV (right).

```
> t.test(Income,PVV)
```

Welch Two Sample t-test

data: Income and PVV

t = 9.02, df = 343.16, p-value < 2.2e-16

alternative hypothesis: true difference in means is not equal to 0

95 percent confidence interval:

3.413664 5.317608

sample estimates:

mean of x mean of y

28.00338 23.63775

#####

```
> wilcox.test(Income,PVV)
```

Wilcoxon rank sum test with continuity correction

data: Income and PVV

W = 31835, p-value < 2.2e-16

alternative hypothesis: true location shift is not equal to 0

#####

```
> t.test(Income-PVV)
```

One Sample t-test

data: Income - PVV

t = 7.6069, df = 206, p-value = 9.833e-13

alternative hypothesis: true mean is not equal to 0

95 percent confidence interval:

3.234163 5.497109

sample estimates:

mean of x

4.365636

#####

```
> wilcox.test(Income-PVV)
```

Wilcoxon signed rank test with continuity correction

data: Income - PVV

V = 16919, p-value = 9.838e-13

alternative hypothesis: true location is not equal to 0

#####

```
> asympcor.test(Income,PVV)
```

Asymptotic correlation test

data: Income and PVV

t = 6.5337 p-value = 6.42e-11

alternative hypothesis: true correlation is not equal to 0

sample estimates:

cor

-0.4541117

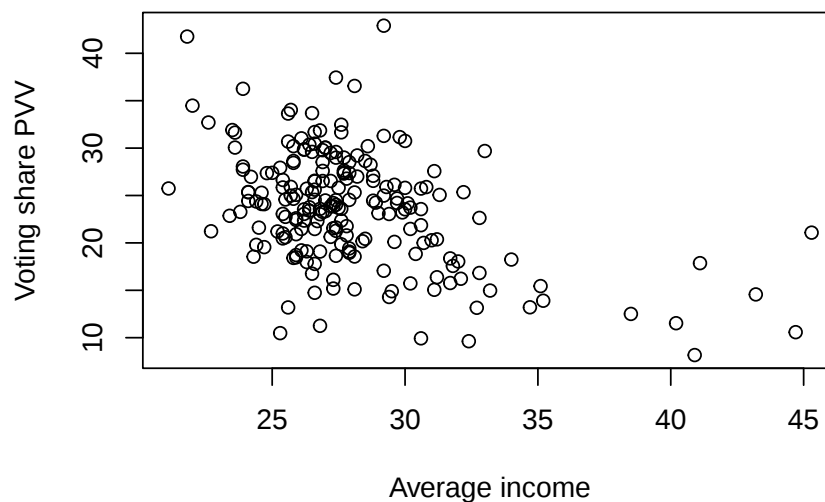


Figure 2: Scatter plot of the variables average income and voting share PVV

#####

```
> chisq.test(Income,PVV)
```

Pearson's Chi-squared test

data: Income and PVV

X-squared = 19251, df = 19158, p-value = 0.3164

Warning:

In chisq.test(Income, PVV) : Chisquared approximation may be incorrect

#####

```
> summary(lm(PVV~Income))
```

Call:

```
lm(formula = PVV ~ Income)
```

Residuals:

Min	1Q	Median	3Q	Max
-15.145	-3.348	-0.301	3.647	20.154

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	44.2310	2.8457	15.543	< 2e-16 ***
Income	-0.7354	0.1008	-7.298	6.31e-12 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 5.292 on 205 degrees of freedom

Multiple R-squared: 0.2062, Adjusted R-squared: 0.2023

F-statistic: 53.26 on 1 and 205 DF, p-value: 6.31e-12

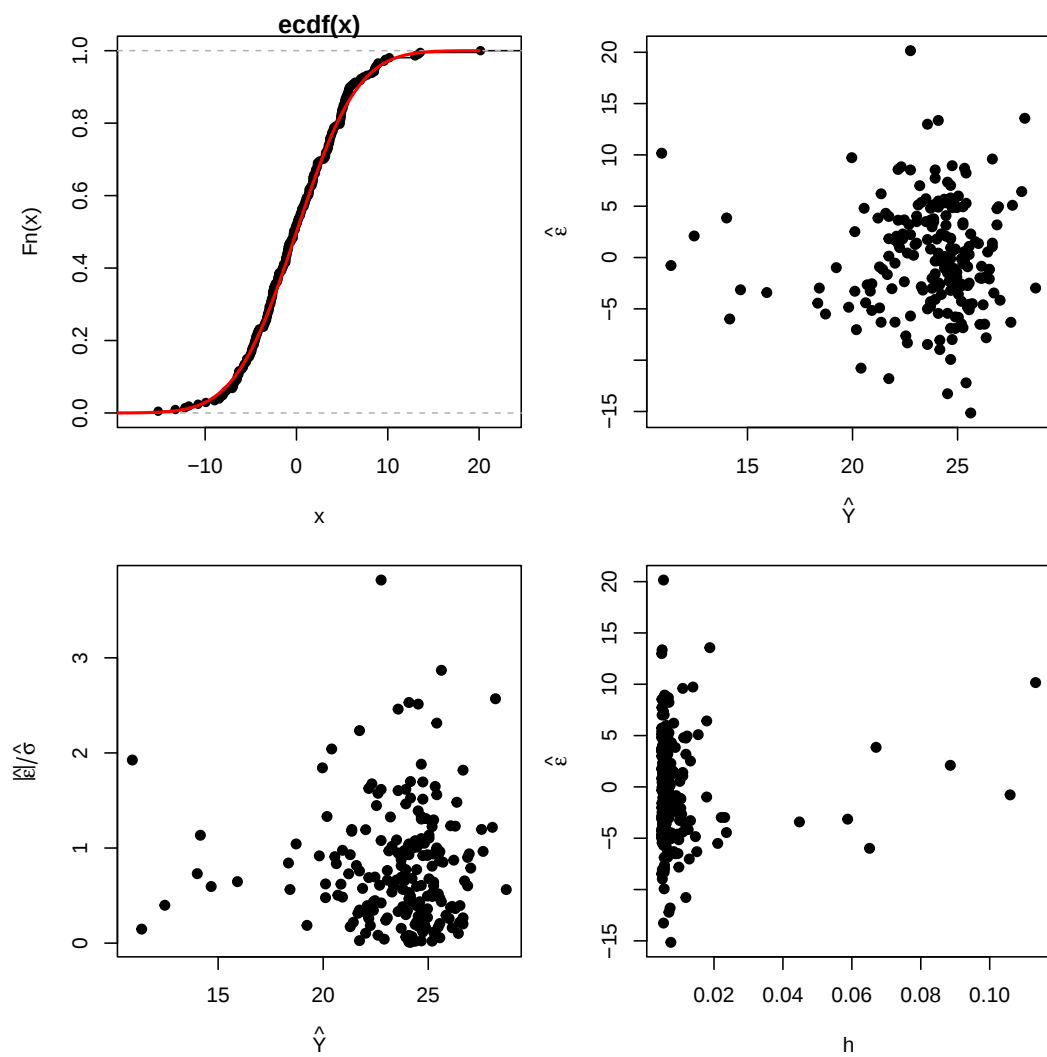


Figure 3: Model assessment for the linear model.